

## Structure Learning: Parameter Estimation II

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## Bayesian Inference

- The MLE is a frequentist inference method. There is another approach to inference called Bayesian inference.
- The key differences between frequentist and Bayesian approaches are shown in the next slides
- See “A primer on Bayesian statistics in Health Economics and Outcomes research” by Anthony O’Hagan and Bryan R. Luce

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## Bayesian Inference

The Nature of Probability

<b>Frequentist</b>	<b>Bayesian</b>
Probability is a limiting, long-run frequency	Probability measures a personal degree of belief
It only applies to events that are (at least in principle) repeatable	It applies to any event or proposition about which we are uncertain

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## Bayesian Inference

The Nature of Parameters

<b>Frequentist</b>	<b>Bayesian</b>
Parameters are not repeatable or random	Parameters are unknown
They are therefore not random variables, but fixed (unknown) quantities	They are therefore random variables

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## Bayesian Inference

The Nature of Inference

Frequentist	Bayesian
Does not (although it appears to) make statements about parameters	Makes direct probability statements about parameters
Interpreted in terms of long-run repetition	Interpreted in terms of evidence from the observed data

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## Bayesian inference

Bayesian inference:

1. Choose probability density  $f(\theta)$  – called the prior distribution that expresses our beliefs about a parameter  $\theta$  before we see any data.
2. We choose a statistical model  $f(x|\theta)$
3. After observing data  $X_1, \dots, X_n$ , we update our beliefs and calculate the posterior distribution  $f(\theta|X_1, \dots, X_n)$

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## Bayesian Inference

Suppose we have  $n$  independent, identically distributed observations  $X_1, \dots, X_n$ . The joint density of the data is:

$$f(x_1, \dots, x_n | \theta) = \prod_{i=1}^n f(x_i | \theta) = L_n(\theta)$$

$$f(\theta | x_1, \dots, x_n) = \frac{f(x_1, \dots, x_n | \theta) f(\theta)}{f(x_1, \dots, x_n)} = \frac{f(x_1, \dots, x_n | \theta) f(\theta)}{\int f(x_1, \dots, x_n | \theta) f(\theta) d\theta}$$

$$= \frac{L_n(\theta) f(\theta)}{\int L_n(\theta) f(\theta) d\theta} = \alpha L_n(\theta) f(\theta)$$

$$\therefore f(\theta | x_1, \dots, x_n) \propto L_n(\theta) f(\theta)$$

Likelihood

Posterior Distribution

Prior (Note: We are not committing to a particular  $\theta$  but using the entire distribution of  $\theta$ )

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## Bayesian Inference

What do you do with the posterior distribution?

- Use the entire distribution (can be clumsy sometimes)
- Get a point estimate by summarizing the center of the posterior – use the mean or mode
- The posterior mean is:

$$\bar{\theta}_n = E[\theta] = \int \theta f(\theta | x_1, \dots, x_n) d\theta = \frac{\int \theta L_n(\theta) f(\theta)}{\int L_n(\theta) f(\theta) d\theta}$$

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## Conjugate Priors

Let's redo the first candy example except this time, we will put a  $Beta(\alpha, \beta)$  prior on  $\theta$ . Recall that  $\theta$  is the probability a candy will be cherry flavored. The posterior has the form:

$$f(\theta | x_1, \dots, x_n) = \frac{f(\theta)L_n(\theta)}{\int f(\theta)L_n(\theta)d\theta}$$

$$f(\theta) = Beta(\alpha, \beta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{\alpha-1}(1-\theta)^{\beta-1}$$

where  $\Gamma(z) = (z-1)!$

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## Conjugate Priors

$$\begin{aligned} f(\theta | x_1, \dots, x_n) &= \frac{f(\theta)L_n(\theta)}{\int f(\theta)L_n(\theta)d\theta} \\ &= \frac{\frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{\alpha-1}(1-\theta)^{\beta-1} \theta^c(1-\theta)^l}{\frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \int \theta^c(1-\theta)^l \theta^{\alpha-1}(1-\theta)^{\beta-1} d\theta} \\ &= \frac{\theta^c(1-\theta)^l \theta^{\alpha-1}(1-\theta)^{\beta-1}}{\int \theta^c(1-\theta)^l \theta^{\alpha-1}(1-\theta)^{\beta-1} d\theta} = \frac{\theta^{c+\alpha-1}(1-\theta)^{l+\beta-1}}{\int \theta^{c+\alpha-1}(1-\theta)^{l+\beta-1} d\theta} \end{aligned}$$

Let's take a look at this term in the denominator

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## Conjugate Priors

Below is the Beta distribution with alpha parameter =  $c + \alpha$  and beta parameter =  $l + \beta$ . Since it is a known pdf, it will integrate to 1.

$$\int Beta(c + \alpha, l + \beta) d\theta = \int \frac{\Gamma(c + \alpha + l + \beta)}{\Gamma(c + \alpha)\Gamma(l + \beta)} \theta^{c+\alpha-1}(1-\theta)^{l+\beta-1} d\theta = 1$$

This is the term in the denominator from the previous page. It is almost a Beta distribution except it is missing the normalization constant in front.

$$\int \theta^{c+\alpha-1}(1-\theta)^{l+\beta-1} d\theta$$

Let's call the normalization constant (the expression with the Gammas)  $c$ . The expression above becomes:

$$\int \theta^{c+\alpha-1}(1-\theta)^{l+\beta-1} d\theta = \frac{1}{c} \int c \theta^{c+\alpha-1}(1-\theta)^{l+\beta-1} d\theta = \frac{1}{c}$$

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## Conjugate Priors

Continuing from where we left off...

$$\begin{aligned} f(\theta | x_1, \dots, x_n) &= \frac{\theta^{c+\alpha-1}(1-\theta)^{l+\beta-1}}{\int \theta^{c+\alpha-1}(1-\theta)^{l+\beta-1} d\theta} \\ &= \frac{\theta^{c+\alpha-1}(1-\theta)^{l+\beta-1}}{\frac{\Gamma(c + \alpha)\Gamma(l + \beta)}{\Gamma(c + \alpha + l + \beta)}} = \frac{\Gamma(c + \alpha + l + \beta)}{\Gamma(c + \alpha)\Gamma(l + \beta)} \theta^{c+\alpha-1}(1-\theta)^{l+\beta-1} \\ &= Beta(c + \alpha, l + \beta) \end{aligned}$$

## Conjugate Priors

- A **conjugate prior** is a family of prior probability distributions with the property that the posterior also belongs to that family.
- eg. the conjugate prior for a Bernoulli is a Beta distribution
- Other useful conjugate priors:

Likelihood	Conjugate Prior	Posterior
Normal	Normal	Normal
Binomial	Beta	Beta
Poisson	Gamma	Gamma
Multinomial	Dirichlet	Dirichlet

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## Conjugate Priors

Why are they useful?

- Since we know the form of the posterior, we can easily calculate statistics such as the mean.
- For example, we know:

$$E[\text{Beta}(\alpha, \beta)] = \frac{\alpha}{\alpha + \beta}$$

- Thus, the mean for the candy example above is:

$$E[\text{Beta}(c + \alpha, l + \beta)] = \frac{\alpha + c}{\alpha + \beta + l + c}$$

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## Conjugate Priors

- You can think of  $\alpha$  and  $\beta$  in the posterior distribution as “virtual counts”
- eg. Using a uniform prior  $\text{Beta}(1, 1)$ , the mean of the posterior becomes:

$$E[\text{Beta}(c+1, l+1)] = \frac{\alpha + c}{\alpha + \beta + l + c} = \frac{1 + c}{2 + l + c}$$

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## Conjugate Priors

$$E[\text{Beta}(c+1, l+1)] = \frac{\alpha + c}{\alpha + \beta + l + c} = \frac{1 + c}{2 + l + c}$$

- If we observe no data, ie.  $c=0, l=0$ , the posterior mean is  $\frac{1}{2}$ , which is what we would expect since we have to pick between the two flavors of lime and cherry
- If we observe lots of data, then the  $c$  term in the numerator and the  $l+c$  term in the denominator dominate the prior

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## Conjugate Priors

- The conjugate prior that is of most relevance to parameter estimation is the Multinomial-Dirichlet
- Recall that a Dirichlet distribution is a generalization of a Beta distribution
- And a Multinomial distribution is a generalization of a Binomial distribution
- If a node in a Bayesian network can take 2 values, the analysis is just like the Beta-Binomial example in previous slides
- If it takes more than 2 values, then you have to use a Multinomial-Dirichlet

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## Conjugate Priors

### Multinomial

$$f(x_1, \dots, x_k | n, p_1, \dots, p_k) = \frac{n!}{x_1! x_2! \dots x_k!} p_1^{x_1} p_2^{x_2} \dots p_k^{x_k}$$

for  $\sum x_i = n, p_i \in [0, 1], \sum p_i = 1$

Note: The parameters  $p_1, \dots, p_k$  from the multinomial are now the random variables in the Dirichlet prior

### Dirichlet

$$f(p_1, \dots, p_k | \alpha_1, \dots, \alpha_k) = \frac{\Gamma(\alpha_1 + \dots + \alpha_k)}{\Gamma(\alpha_1) \dots \Gamma(\alpha_k)} p_1^{\alpha_1 - 1} \dots p_k^{\alpha_k - 1}$$

for  $p_i \geq 0, \sum p_i = 1$

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## Conjugate Priors

Likelihood	Conjugate Prior	Posterior
Binomial( $x   n, p$ )	Beta( $\alpha, \beta$ )	Beta( $x+\alpha, n-x+\beta$ )
Multinomial( $x_1, \dots, x_k   n, p_1, \dots, p_k$ )	Dirichlet( $p_1, \dots, p_k   \alpha_1, \dots, \alpha_k$ )	Dirichlet( $x_1 + \alpha_1, \dots, x_k + \alpha_k$ )

For Beta-Binomial posterior:  $E[p] = \frac{x + \alpha}{n + \alpha + \beta}$

For Dirichlet-Multinomial posterior:  $E[p_i] = \frac{x_i + \alpha_i}{n + \sum_j \alpha_j}$

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## Conjugate Priors

Suppose you were asked to estimate  $P(\text{Price} = \text{Low} | \text{Type} = \text{Sedan}, \text{Color} = \text{Silver})$ .

Notice that this distribution is a multinomial distribution with  $n = 2$  (because there are 2 records with Color=Silver, Type=Sedan) and  $p_{\text{low}}, p_{\text{medium}}, p_{\text{high}}$  corresponding to when Price is low, medium, and high.

Now suppose I tell you to use a Dirichlet prior where all the  $\alpha_i$  are 1.

Color	Type	Price
Silver	Sedan	Low
Black	Sedan	Medium
Silver	Pickup	High
Silver	Sedan	Low
Red	SUV	High

Estimate  $P(\text{Price} = \text{Low} | \text{Color} = \text{Silver}, \text{Type} = \text{Sedan})$

$$= \frac{\#(\text{Color} = \text{Silver AND Type} = \text{Sedan AND Price} = \text{Low}) + 1}{\#(\text{Color} = \text{Silver AND Type} = \text{Sedan}) + 3}$$

$$= \frac{2 + 1}{2 + 3} = \frac{3}{5}$$

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